

# DEMOCRACY, MARKET LIBERALIZATION, AND POLITICAL PREFERENCES

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*Abstract*—We estimate the impact of market development and democratization on subjective political preferences. We rely on the specific situation of frontier zones and the considerable regional variations in culture and economic development in the countries of the former socialist bloc for identification. Using a survey conducted in 2006, we find a positive and significant effect of democracy on support for a market economy, but no effect of market liberalization on support for democracy. Hence, in contrast with the conventional wisdom concerning the sequencing of political and economic reforms, democratization may become a necessary condition to obtain public support for further economic liberalization.

## I. Introduction

THIS paper questions the conventional wisdom concerning the sequencing of political and economic reforms in developing countries. A popular idea is that market liberalization should come before democratization, in the early stage of a country's development. Supporters of this strategy point to the erratic reform path experienced by some Central and Eastern European countries of the former socialist bloc, which predominantly chose rapid and simultaneous political and economic liberalization in the 1990s (Dewatripont & Roland, 1992; Roland & Verdier, 2003; Godoy & Stiglitz, 2006). Another example is Latin America, where pervasive economic crises seem to illustrate how democracy can be an obstacle to the development of the market when leaders need to impose unpopular reforms while being responsible to their constituencies. By contrast, in China, the continued hold of the Communist party over political power is taken to be a key ingredient in the spectacular development of the market economy. In this view, the optimal route is to develop market institutions in a first stage of development and consider democratization at a later stage.

Pushing the argument one step further, some authors have argued that the desire for political freedom and democratic institutions does not arise until countries reach a certain degree of material comfort and market liberalization (Lipset, 1959; Miller, Hesli, & Reisinger, 1994; Miller, Reisinger, & Hesli, 1996). Hence, the argument goes, not only is it more practical to postpone democracy until the market is developed, but this sequence also meets citizens' preferences.

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The conventional wisdom concerning the linkages between political and economic systems has changed considerably over time. The idea that "modern democracy is a by-product of the capitalist process" (Schumpeter, 1942) finds its roots in the writings of Montesquieu, Steuart, Hume, and many philosophers of the eighteenth century, who saw trade as a quiet passion (*le doux commerce*) and hypothesized an affinity between trade and social harmony. Skepticism about this spontaneous "psychomachia" (Hirschman, 1977) started with Smith and de Tocqueville and became radical in the nineteenth century (see Marx, 1990, or Mill, 1951). Today the consensus seems to be that "development" implies both market liberalization and political democracy,<sup>1</sup> but with the market coming first. Lipset (1959, p. 80), for example claims that "industrialization, urbanization, high educational standards and a steady increase in the overall wealth of society [are] basic conditions sustaining democracy." The similarity between market competition and political competition between organized groups has been used to underscore the economic preconditions of democracy (Dahl, 1982; Andrain, 1984). Another line of reasoning is that postponing political liberalization gives more leeway to reformers when reform measures potentially hurt important groups of the population. This argument builds on the political economy of transition literature (Aslund, Boone, & Johnson, 2001; Dewatripont & Roland, 1992; Roland, 2001; Roland & Verdier, 2003), which discusses the relative pros and cons of democracy versus authoritarianism from the point of view of facilitating economic reforms.

Beyond these theoretical approaches, the existing empirical literature does not offer a clear picture in terms of optimal sequencing of political and economic reforms. A vast literature focuses on the relationship between income growth and democracy. Most papers conclude with a positive relation running from income to democracy (Przeworski, 2004; Boix, 2003; Boix & Stokes, 2003; Epstein et al., 2006; Papaioannou & Siourounis, 2008; Londregan & Poole, 1996; Minier, 2001; Burkhart & Lewis-Beck, 1994). However, Acemoglu et al. (2005a, 2005b; Acemoglu, Bautista et al., 2007), show that this relationship is not robust to the inclusion of country fixed effects, so there is "no evidence that economic development has a causal effect on democracy" (Robinson, 2006). Concerning the reverse relationship, from political regimes to economic growth, "the findings are all over the place" (Persson & Tabellini, 2008). Many papers find no impact of democracy (or democratic transitions) on income growth (Londregan & Poole, 1996; Burkhart & Lewis-Beck, 1994; Helliwell, 1994; Barro, 1996; Rodrik, 2000), but some do find a posi-

<sup>1</sup> Hence the concept of the "end of history" (Fukuyama, 1992).

tive relation (Leblang, 1997; Minier, 1998; Persson & Tabellini, 2006, 2007, 2008; Rodrik & Wacziarg, 2005; Tavares & Wacziarg, 2001). A meta-analysis by Przeworski and Limongi (1993) confirms that the literature is inconclusive. These authors reviewed 18 studies and 21 findings: a third of these conclude that democracy is superior in promoting economic growth, another third claims that authoritarian regimes are better at promoting growth, and the last third finds no significant relationship.

Concerning the relation between democracy and market liberalization, the literature is less abundant. Persson and Tabellini (2006) find that democracy is associated with economic reforms, while economic liberalization only “sometimes” leads to democratization, but that liberalizing the economy first is better for subsequent GDP growth. Giavazzi and Tabellini (2005) claim that causality is more likely to run from political to economic liberalization. A series of papers exploits the experience of postsocialist countries, often hinging on election data or individual-level surveys: they essentially conclude that causality link runs from democracy to market liberalization (Fidrmuc, 2003) or from democracy to the support for market liberalization (Dethier, Ghanem, & Zoli, 1999; Hayo, 2004).

In this literature, the effort dedicated to the identification strategy is of uneven importance, consisting of instrumentation with lagged variables (Dethier et al., 1999), country fixed effects (Papaioannou & Siourounis 2008; Londregan & Poole 1996), country and year fixed effects (Persson & Tabellini, 2006; Acemoglu et al., 2005a, 2005b; Acemoglu, Johnson, et al., 2007), simultaneous equations (Tavares & Wacziarg, 2001), Granger tests (Fidrmuc, 2003), difference in difference, and matching (Persson & Tabellini, 2007). However, Acemoglu et al. (2005a, 2007), Acemoglu, Johnson et al. (2005) and Acemoglu, Johnson, and Robinson (2001), convincingly argue for the quasi impossibility of excluding the influence of omitted variables such as culture, religion, or colonial institutions that preside over the evolution of all the magnitudes of interest. The influence of such distant past “critical junctures” precludes establishing a univocal direction of causality. This identification problem is also contained in the very idea of modernization theory that the same development dynamics favor both democracy and market development (Lipset, 1959). Even papers based on panel data with time and country fixed effects are likely to suffer from omitted variables, as the relevant dimension of heterogeneity is not necessarily national: it can be more local (for example, eastern versus western Ukraine) or wider than the nation (the lasting influence of past empires).

In this paper, we try to overcome this simultaneity bias by using the precise spatial information available in our data set. We exploit a new set of micro evidence from an original survey of 28 transition economies, the Life in Transition Survey, which was implemented in summer 2006 (see EBRD, 2007). This survey offers the possibility of identifying precisely the geographical location of the pri-

mary sampling units. We then hinge exploit the national and regional sources of heterogeneity.

We first estimate the causal relationship that runs from actual democratization to popular support for a market economy. Our empirical identification strategy consists of a spatial regression discontinuity design that relies on the specificities of frontier zones. Our main identification assumption is that people who live on either side of an integrated frontier zone experience different political regimes but share the same experience of the market and, often, the same historically inherited cultural attitudes toward the market and democracy on both sides of the frontier. This assumption is particularly plausible the often artificial frontiers of the former USSR and other formerly integrated regions such as the Ottoman or the Habsburg empires. This assumption is tantamount to keeping constant the omitted variables that usually bias estimations of the relationship between market development and the support for democracy.

We also estimate the reciprocal relationship: from actual market development to popular support for democracy. Here, we exploit within-country regional variations. We rely on the fact that the degree of market development is heterogeneous across different regions of the countries included in the survey, whereas people who live in the same country share a common experience of democracy. We build an index of regional market development that reflects the share of the modern sector of the economy, composed of private, small, and medium-size firms. We then compare the support for democracy in the various regions of a given country, where market development varies but political and democratic rights are similar.<sup>2</sup>

We do not pretend to explain the long-run causality between democracy and market development; instead, we ask whether the demand for democracy emerges after a certain degree of market development is reached and whether, conversely, democratization is more likely to be an obstacle or an ingredient to citizens’ support for market liberalization.

The main result of the paper is that democracy increases popular support for the market, whereas, contrasting with the conventional wisdom, economic liberalization does not clearly enhance the support for democracy. This result is robust to alternative indices of democratization and market liberalization, as well as to other proxies of preferences for democracy.

Section II describes the data. Section III presents the identification strategies for the two relationships: between democracy and support for the market, and between market development and support for democracy. Section IV discusses the empirical results. Section V concludes.

<sup>2</sup> The assumption of similar political rights within a country may be questioned, but as we detail later in the paper, most countries in our sample are highly centralized politically. Even in the case of the Russian Federation, the strong presidential regime helps to attenuate the (statistical) problem.

## II. Data

Our study uses the Life in Transition Survey (LITS) conducted by the European Bank for Reconstruction and Development and the World Bank in 2006 in 28 post-transition countries.<sup>3</sup> Respondents to the survey were drawn randomly, using a two-stage sampling method, with census enumeration areas as primary sampling units (PSUs)<sup>4</sup> and households as secondary sampling units. This nationally representative survey includes 1,000 observations per country, making a total of 28,000 observations.

### A. Support for the Market and for Democracy

Support for the market is analyzed using the following question:

Which of the following statements do you agree with most?

- A. A market economy is preferable to any other form of economic system.
- B. Under some circumstances, a planned economy may be preferable to a market economy.
- C. For people like me, it does not matter whether the economic system is organized as a market economy or as a planned economy.

We analyze the probability of choosing one of the first two statements.

Concerning the support for democracy, we analyze the probability of choosing one of the first two proposed answers to the question:

Which of the following statements do you agree with most?

- A. Democracy is preferable to any other form of political system.
- B. Under some circumstances, an authoritarian government may be preferable to a democratic one.
- C. For people like me, it does not matter whether a government is democratic or authoritarian.

One concern with this type of subjective questions is that answers might reflect some personality traits rather than genuine preferences. Although there is no way to fully eliminate this concern with cross-section data, we control for several observable individual characteristics that are likely to be correlated with unobserved individual heterogeneity. We also checked that people do not answer these two questions in a systematic way. For example, people who declare that “under some circumstances, a planned economy may be preferable to a market economy” are equally

likely to choose the first or the second answer to the question about democracy.

We also use an alternative measure of support for democracy based on the following survey question:

To what extent do you agree that the following are important for your country?

- A. Free and fair elections
- B. Law and order
- C. Freedom of speech
- D. A press that is independent from the government
- E. A strong political opposition
- F. A courts system that defends individual rights against abuse by the state
- G. A courts system that treats all citizens equally, rather than favoring some over others
- H. Protection of minority rights (religious, ethnics, etc. . .)
- I. Freedom to travel abroad

All of the questions used this scale: strongly disagree, disagree, neither agree nor disagree, agree, strongly agree, difficult to say).

We construct an index measuring “attachment to democracy,” which sums the number of times a person “agrees” or “strongly agrees” that the items listed are important. The regional index varies from 0 to 9, with an average of 4.5. The average score is highest in Montenegro and Croatia (7) and lowest in Tajikistan (3).

### B. Frontier Zones

The LITS is based on primary sampling units (PSUs), each containing twenty observations (surveyed persons). Information on the precise location of these PSUs is available in a series of country maps and through the location names, which we use in order to identify groups of PSUs that are located in the immediate vicinity of a political frontier. We define frontier zones as groups of PSUs that are quasi-adjacent and located on both sides of a frontier (less than 25 kilometers from the border).

As discussed in section III, the validity of our identification assumption relies on the level of trade openness and market integration across the borders. We thus distinguish open frontiers from those that are closed or restricted because of political tensions and disputed territories or geographical obstacles. We hence exclude the (closed) frontiers between Armenia and Azerbaijan, Georgia and Russia, Moldova and Ukraine, and all Uzbek borders. We keep all the other (open) frontiers. We are left with 36 frontier zones, that contain between 40 and 460 observations, concentrated in 2 (Slovak Republic-Ukraine) to 24 (Croatia-Slovenia) PSUs.<sup>5</sup>

<sup>3</sup> Turkmenistan was not included in the survey; neither was Kosovo.

<sup>4</sup> PSUs were selected randomly, with probability proportional to size.

<sup>5</sup> For a complete list of frontier zones considered in the analysis, please refer to the working paper version of this text (Grosjean & Senik, 2007).

In the case of the ten new EU members, persons and goods are free to circulate across borders.<sup>6</sup> This is also true of most part of Central, Eastern and Southeastern Europe (for example, Albania, Macedonia, and Montenegro;<sup>7</sup> Bulgaria and Macedonia or Moldova and Romania). Many countries of the survey are integrated in Euroregions, which are “arrangements for cooperation between units of local or regional government across borders with the purpose of promoting common interests and enhancing the living standards of the border populations.”<sup>8</sup> Two CIS countries, Belarus and Ukraine, are also integrated into Euroregions.<sup>9</sup> These countries are also largely integrated with Russia, not only historically but also formally, in the framework of the Neman Euroregion which also includes Lithuania and Poland. Finally, twelve countries of the former USSR are members of the Commonwealth of Independent States (CIS), created in December 1991 in order to promote regional economic integration.<sup>10</sup> Several other regional agreements were signed among members of the CIS, among which the creation of the Eurasian Economic Community,<sup>11</sup> or the creation of a Common Economic Space including Belarus, Kazakhstan, Russia and Ukraine.

### C. Scores of Democratization

We principally hinge on the Freedom House democracy score (Freedom House, 2006a), which averages national ratings for the electoral process, civil society, independent media, independence of the judicial system, and corruption. This score takes values from 1 to 7, with 1 representing the highest level of democratic progress and 7 the lowest. We have recoded it in order to present the score of democracy in an ascending order. In an alternative specification, we dichotomize this score, and for each pair of adjacent countries, we build a dummy variable that takes a value of 1 in the PSUs located in the most democratic country of each pair and 0 in PSUs located in the less democratic country (according to the Freedom House indicator).

For robustness, we verify that our results hold when we use other political scales, such as Freedom in the World (Freedom House, 2006b), BTI indicators (Bertelsmann Stiftung, 2005), or Polity IV (CIDCM, 2006). We favor the Freedom House democracy score over other ratings,

<sup>6</sup> Bulgaria, the Czech Republic, Estonia, Latvia, Lithuania, Hungary, Poland, Romania, the Slovak Republic, and Slovenia.

<sup>7</sup> As well as Kosovo, but Kosovo was excluded from our sample.

<sup>8</sup> Euroregions are formed in order to address issues that involve important local cross-border externalities, hence the need for regional cooperation. See [www.coe.int](http://www.coe.int). Albania, Bosnia and Herzegovina, Croatia, Montenegro, and Slovenia are part of the Adriatic Euroregion; Latvia, Lithuania, Poland, and Russia are part of the Baltic Euroregion (along with Sweden and Denmark); Hungary, Romania, and Serbia are integrated in the Danube-Kris-Mures-Tisza Euroregion

<sup>9</sup> The Białowieża Forest Euroregion includes Poland and Belarus, and the BUG Euroregion includes Belarus, Poland, and Ukraine.

<sup>10</sup> Azerbaijan, Armenia, Belarus, Georgia, Kazakhstan, Kyrgyzstan, Moldova, Russia, Tajikistan, Turkmenistan, Uzbekistan, and Ukraine.

<sup>11</sup> Belarus, Kazakhstan, Kyrgyzstan, Russia, Tajikistan.

because it is entirely focused on democratic rights, as opposed to Freedom in the World or BTI, which include ratings on civil and economic rights, and because it is more contrasted than other indicators, including Polity IV.<sup>12</sup>

### D. Indices of Market Liberalization

Using questions about the respondents’ first, second and third jobs, we build a regional score of market liberalization.<sup>13</sup> The score is the regional proportion of respondents in the active labor force who are self-employed with more than five employees or have a formal labor contract and work in a small enterprise, work in a medium enterprise, work in a private firm, or work in a newly created enterprise (since 1989).<sup>14</sup>

During the Soviet era, socialist economies were distinguished by their exceptionally low proportion of small and medium-sized enterprises (SMEs). The logic of central planning favored the organization of production and distribution in large units. We thus interpret the presence of SMEs as an indicator of market development. The presence of private and of newly created firms is also a sign of progress in the transition toward a market economy, an essential aspect of which is the process of privatization of the formerly predominant state-owned sector and the elimination of former monopolies under the pressure of new competitive firms (Berkowitz & Jackson, 2005). Self-employed persons with at least five employees are also part of this new economic pattern that is typical of a market economy and was absent from the landscape of planned economies. We do not count self-employed persons with fewer than five employees on the grounds that those are less likely to be small firms than forms of quasi-unofficial economy or what is sometimes called “disguised unemployment” (Earle & Sakova, 2000). Finally, we interpret the existence of an employment contract as a sign that the person is working in the official sector rather than in the informal sector, an indication of development of the market economy.

This index varies from 0 to 5 with an average of 2.37. It is highest, on average, in Latvia (2.78), followed by the two other Baltic States, and is lowest in Azerbaijan (1.90) and Belarus (1.92).

One concern about our index of market liberalization is that its quality is limited by the lack of representativeness of the data at the regional level. De facto, the average number of observations by region is 82. However, there is no

<sup>12</sup> In particular, it provides a ranking for Bosnia and Herzegovina, which Polity IV does not. Furthermore, several countries have a similar Polity IV. As a result, results lose some degree of significance when Polity IV is used.

<sup>13</sup> Multiple jobs are frequent in transition countries.

<sup>14</sup> This index thus excludes workers in state-run enterprises or collective farms, enterprises that were already existing in 1989, large enterprises of more than 100 employees, self-employed with strictly fewer than five employees, and employees without a former labor contract. Our results are robust to variations of that index, for example, by considering workers in firms larger than 100 employees or by ignoring whether employees have a formal labor contract.

alternative regional index of private sector development or market liberalization that is comparable across countries of our sample. We also regret that indices of industrial concentration are not available at the regional level for the whole set of countries in the sample.<sup>15</sup>

As a first validation of our index, we compared it to other available indices of market liberalization and privatization at the country level. We used several EBRD transition indicators and constructed a transition indicator as the cumulated score of large-scale privatization, small-scale privatization, enterprise restructuring, price liberalization, competition policy, banking reform, and interest rate liberalization. The correlation coefficient between this score and our index of market development is 0.77.

As an additional test of validity, we confronted our index of market development with regional indices of the shares of private, public, collective, or foreign firms in Moldova in 2006, based on firm census data from the World Bank. To remain as close as possible to the LITS index, we considered only private firms between 5 and 100 employees. The correlation coefficient between our LITS-based index and the Moldovan index is 0.34 (significant at 1%). Note that the LITS-based index considers households' addresses, while the Moldovan census records firms' registered address, which is not necessarily the same address where they operate. The LITS-based index also excludes firms that existed before 1989, while the Moldovan data do not contain such information. This could explain why the correlation coefficient between the two indices is not higher.

All descriptive statistics are presented in appendix.

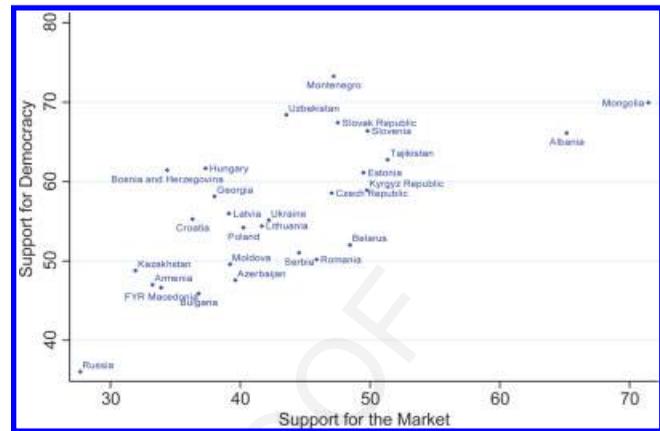
### III. Identification Strategy: Spatial Regression Discontinuity

In order to discern the direction of causality between market and democracy, one should ideally rely on a situation in which one variable is exogenously frozen while the other randomly takes different values across countries. Of course, this ideal setting could never exist. On the contrary, market liberalization and democracy are processes that follow highly intertwined dynamic evolutions and depend on countries' historical background.

Even in the case of transition countries, where both democracy and the market have been abolished by the communist experience for at least half a century, these institutions have evolved in parallel since 1989, probably under the influence of common factors, such as popular pressure and the influence of European and international organizations.

<sup>15</sup> In our view, indices of this type would be better suited, if they were available, than some often used indicators based on governance, the protection of legal rights, the protection of minority shareholders, or indices of price liberalization. The last have two important drawbacks: they are available only at the national level and often reflect the progress of the rule of law, that is, of democracy itself rather than that of the freedom of transactions on the market.

FIGURE 1.—SUPPORT FOR THE MARKET AND FOR DEMOCRACY



Percentage of Respondents Who Favor the Market or Democracy  
Percentage of respondents who favor the market or democracy, that is, who declare, "A market economy is preferable to any other form of economic system" and "Democracy is preferable to any other form of political system" (see section IIA).

As an illustration, figure 1 shows the strong general cross-country relationship between the average support for the market and the average support for democracy in the 28 countries covered by the Life in Transition Survey. Regional differences are also visible. In particular, countries of Central and Eastern Europe and the Baltic states, which are most advanced on the road to a free market and a full-blown democracy, are also the most supportive of the two processes.

Obviously, using the pooled cross-section data of LITS and running a naive regression of support for democracy on an index of market development, or of support for the market on a democratic index, would come across serious identification problems. The observed correlation would not readily be interpretable in terms of causality because of the influence of omitted variables that affect market development, democracy, and the subjective support for either political democracy or a market economy. We propose two different identification strategies in order to isolate the direction of causality from democracy to support for a market economy and from market development to support for democracy.

#### A. Democracy and Support for the Market

The first question is whether a higher degree of democracy is likely to influence the support for a market economy. To address this question, one needs to overcome the problem that subjective support for a market economy may be due to both the degree of democracy and the degree of market development itself, both variables being likely to evolve at a parallel pace. It can also be suspected that common cultural factors influence attitudes toward both the market and democracy. This potential heterogeneity may come from both national and more local regional history. In other words, one would like to estimate the naive equation (1) of the support for the market of individual  $i$  in region  $r$  of

country  $j$  depending on the degree of democracy in country  $j$  ( $Democracy_j$ ):

$$SupportMarket_{irj} = a_0 + a_1Democracy_j + a_2X_{irj} + u_i, \quad (1)$$

but one suspects that the true relationship is (1')

$$SupportMarket_{irj} = a_0 + a_1Democracy_j + a_2Market_{rj} + a_3Culture_{rj} + a_4X_{irj} + u_i, \quad (1')$$

where subscripts  $i$  stand for individuals,  $r$  for regions, and  $j$  for countries;  $X_{irj}$  stands for the socio-demographic characteristics of respondent  $i$  living in region  $r$  of country  $j$ ,  $Market_{rj}$  reflects the extent of market development in region  $r$ ;  $Culture_{rj}$  embodies regional or national cultural factors; and  $u_i$  is the error term.

Our strategy consists of trying to keep the second and third terms of equation (1') constant. As our analysis is based on individual data, we need to find individuals who, in an exogenous way, are confronted with different levels of democracy but to the same degree of market development and who share the same inherited culture regarding the politicoeconomic system.

*The Specificity of Frontier Zones:* We rely on the fact that the survey covers a large continuous territory comprising many contiguous countries, including observations at the political border between pairs of countries. The idea is to identify observations in frontier zones. We assume that because of spatial integration, people who live in open frontier zones are influenced by common cultural values and economic conditions, even though they live on different sides of the frontier. In other words, in frontier zones, while political institutions and the level of democracy vary across the frontier, there is local invariance in the level of market development. This relies on two types of arguments.

The first one is the well-documented high level of interregional trade in frontier zones (Feenstra, 2004). It is well known that in frontier zones, whenever it is possible, people do not hesitate to cross the border to work or shop. Seasonal migration (Gould, 1994) and "shuttle trade" (see Grafe, Raiser, & Sakatsume, 2005 in the case of Central Asia) are specific forms of transactions that are common place near the border, based on local social and business networks (Rauch, 2002). Hence, regional integration is a fact of everyday life and certainly influences the perception of the market by residents of the frontier zone.

To be sure, political frontiers also act as a barrier to trade between countries, as attested by the frontier effect familiar to trade economists (McCallum, 1995; Wolf, 2000). However, recent papers have shown that the lower intensity of trade across frontiers is essentially due to the size of countries, pricing effects (Anderson & van Wincoop 2003), language differences (Melitz, 2002), and currency heterogeneity (Parsley & Wei, 2001; Frankel & Rose, 2002), rather than

to political discontinuity. It is likely that these factors are less important for inhabitants of a frontier zone than for the average citizen of a country, as the former usually speak at least one common language and are certainly less reluctant to hold the currency of the other country, which they can use frequently. Hence, although the volume of international transactions is lower than that of intracountry transactions, bilateral trade between two countries is likely to be particularly vivid in the vicinity of the border between them. This is consistent with the well-established finding, using gravity models, for instance, that distance is an obstacle to trade, in particular, internal distance to the border (Melitz, 2007).

Second, in the specific case of the fifteen former Soviet republics, regional integration was a hard fact until the early 1990s. Under the socialist system, the economy of the Soviet republics was under the control of the centralized organization of material resources by the Soviet plan based in Moscow. Many countries, particularly in Central Asia and the Caucasus, have inherited from the Soviet Union integrated infrastructure networks, a positive factor of trade and regional integration (Broadman, 2006). Fidrmuc and Fidrmuc (2003) show that trade among the former constituent republics of the Soviet Union, of Yugoslavia, and of Czechoslovakia has remained well above any definition of normal trade intensity, even after the disintegration of the three federations. In the case of Central Asia and Caucasus, the very slow change in patterns of trade since 1990 (Babetskii, Koukhartcouk, & Raiser, 2003; Broadman, 2006) has led researchers to argue that these countries are "overtrading" among themselves (Grafe et al., 2005).<sup>16</sup>

Anecdotal evidence about the vivid economic activity in frontier zones of the former socialist bloc can be found in Holtom (2006), in the particular example of the Russian-Lithuanian border zone: "Every day the population in Russia-Lithuania border town of Sovetsk increases by several thousand thanks to Lithuanian citizens who every morning cross the Koroleva-Luiza bridge on their way to work in the town" . . . 12 Lithuanian factories have opened in Sovetsk since 2000, providing employment to the people of the town and of the neighboring Lithuanian regions. Similarly, the opportunities for trade along the Polish-Ukrainian-Belarusian borders gave rise, after 1993, to bazaars and open air markets along Poland's eastern border. Based on these arguments and observations, we thus assume that inhabitants of certain frontier-zones share the same experience of the market, even when they live in different countries with different political institutions.

We verify that the degree of market development is more similar between two adjacent frontier zones than it is on average between two adjacent countries. Using the index of

<sup>16</sup> Using relative prices of a bundle of goods to complement official trade data, Grafe et al. (2005) show that the impact of borders on trade between Kazakhstan, Kyrgyzstan, Uzbekistan, and Tajikistan is much smaller than what the view of cumbersome crossing border procedures and licensing systems would imply. The authors attribute this result to the large development of shuttle trade in this region.

market development as defined in section II D, we find that on average, the correlation between indices of market development is three times as high between adjacent frontier zones of the sample ( $z_i$  and  $z_j$ ) as it is between adjacent countries ( $i$  and  $j$ ) of the sample (see table A2 in Appendix). If we restrict the analysis to subsets of formerly more integrated countries, the proximity between adjacent frontier zones appears even higher. For instance, in Central Asia, the correlation between two adjacent frontier zones is 0.78 against 0.11 in two adjacent countries (row 5 of table A1). We also checked that in frontier zones, the structure of employment in terms of occupations and sectors is not statistically different on either side of the borders (table A3). By contrast, this was not true of entire adjacent countries.

*Cultural Zones:* Market integration at frontier zones is useful to eliminate the risk that the measure of support for market liberalization reflects the actual market development, but what about cultural omitted factors? A first element of answer is that the citizens of the former Socialist bloc, and in particular the Soviet Union, have been living for 45 to 70 years in a common political system marked by strong official ideological values concerning the market. It is not far-fetched to assume that they share a common heritage in terms of attitudes toward the market (Alesina & Fuchs-Schundeln, 2007). The very idea of national culture is that countries' past experience continues to exert long-term effects (Fernandez & Fogli, 2005).<sup>17</sup> We therefore rely on the idea that citizens of countries that have belonged to formerly highly integrated zones share a common culture, that is, common inherited attitudes toward the market and democracy. There are some subsets of the transition countries in which this assumption is particularly appealing. Regions that belonged to the Ottoman Empire (Albania, Bosnia and Herzegovina, Bulgaria, Bessarabia, Crimea, FYR Macedonia [FYROM], Moldavia, outer Montenegro, Serbia except Vojvodina, and Wallachia) developed under the same rule for several centuries (see figure A1).<sup>18</sup> The same is true of regions of the Habsburg Empire (Croatia except Dalmatia, Czech Republic, Hungary, Polish Silesia, Slovak Republic, Slovenia, Transylvania, and Vojvodina), of Prussia (Estonia, Latvia, Polish Silesia, Pomerania, Royal Prussia), countries of the former Polish-Lithuanian Commonwealth (which included Poland, Ukraine, Lithuania, Belarus, and western parts of Russia), countries of the USSR (1922–1991), or countries of the Former Yugoslavia (Bosnia and Herzegovina, Croatia, FYROM, Montenegro, Serbia, and Slovenia), which shared the same rule for several decades (1918–1991). Countries of Central Asia have also shared common influences practically until their independence in the early

1990s, starting with Alexander the Great's Empire, then under the Persian, Turkish, Mongol, and Timuride empires and finally under the Russian Empire and the Soviet Union. We thus retain these cultural and historical groupings in order to deal with the potential impact of cultural factors on attitudes toward the politicoeconomic system (we run subregressions inside each of these cultural zones).

We also rely on the fact that current frontiers of many transition countries are more or less artificial divisions of formerly integrated jurisdictions and have been frequently changing. The idea that frontiers are arbitrary has been challenged by recent empirical work; for example, Schulze and Wolf (2007) show that frontiers can emerge spontaneously in a manner that separates heterogeneous populations and include together more homogeneous groups (see also Alesina & Spolaore, 1997, or Alesina, Spolaore, & Wacziarg, 2000). In the case of the former USSR, however, frontiers and entire ethnic groups have, on the contrary, been artificially displaced under the Stalinist regime, mixing heterogeneous populations together and tracing arbitrary frontiers in order to minimize the risk of nationalism, secessionist movements, and mobilization against the Soviet regime (Werth, 2001; Cadot, 2005). Of course, after the dissolution of the Soviet Union, new frontiers, not all of them arbitrary, were traced. However, the legacy of the Stalinist territorial policy made it impossible to delimit ethnically homogeneous territories. The fact that people were free to move after the delimitation of the border does not alter our main assumption that they share many of the politicoeconomic attitudes with the inhabitants of the other frontier zone. Even in Central and Eastern Europe, it was only after World War I and the collapse of the Ottoman, Czarist, Prussian, and Habsburg empires that nation-states became established. From the start, questions lingered over their viability, precisely because of unresolved national questions that did not correspond to the borders of these new states, but on the contrary widely cut across them (Batt, 2002).

*Estimating support for the market in frontier zones:* Our test therefore consists, for the set of inhabitants of open frontier zones, of regressing individual support for the market on an index of democracy, controlling for frontier zone dummies and other sociodemographic characteristics of respondents. The assumptions of market integration and common culture at the frontier between two formerly integrated countries imply that the third and fourth terms of equation (1') are constant and can thus be dropped out. We thus estimate equation (1'') on the subsamples of frontier zones. We regress the support for the market of individual  $i$  living in frontier zone  $k$  of country  $j$ , on the level of democracy of country  $j$ :

$$\text{Support Market}_{ikj} = a_0 + a_1 \text{Democracy}_j + a_2 X_{ikj} + a_3 C_k + u_i, \quad (1'')$$

where  $\text{Democracy}_j$  corresponds to the Freedom House democracy score of country  $j$ ,  $X_{ikj}$  stands for the sociodemo-

<sup>17</sup> In Bisin and Verdier (2000) and Fernandez and Fogli (2005) define culture as long-term inertia in preferences.

<sup>18</sup> Of course, the delimitation of the former Habsburg, Prussian, or Ottoman empires has varied greatly along history. We retain in this paper only regions that have belonged to these former Empires for at least 200 years (Sheperd, 1911; EurAtlas, 2008).

graphic characteristics of respondent  $i$ , living in country  $j$  and frontier zone  $k$ ,  $C_k$  is a vector of frontier zone dummies, and  $u_i$  is the error term. In an alternative specification, we run the same regression on a dummy variable indicating whether the country of residence of an individual is more democratic than the adjacent country (we dichotomize the variable measuring the score of democracy). We also check that our results hold when alternative measures of democracy are used (see section IVA).

#### B. Market Development and Support for Democracy

To isolate the causation running from market liberalization to support for democracy, we need to overcome the symmetrical problem, that is, avoid the contamination of the actual degree of democracy and the influence of cultural factors on support for democracy. This implies keeping constant the second and third terms of equation (2):

$$\begin{aligned} \text{SupportDemocracy}_{ij} = & b_0 + b_1 \text{Market}_{rj} + b_2 \text{Democracy}_j \\ & + b_3 \text{culture}_{rj} + b_4 X_{ij} + u_i. \end{aligned} \quad (2)$$

Here, we rely on the considerable within-country regional variations in terms of market development (Zhuravskaya, 2006; EBRD, 2006). We build an index that reflects the regional emergence of private, small, and medium enterprises and the formal sector (see section IID).

We start with the simplifying assumption that political institutions and culture are the same in a given country. We consider that as far as democracy is concerned, the changes mainly operate at the national level, whereas the emergence of the market is more differentiated across regions. Of course, in some countries, the degree of democracy may vary across regions if local governments have an important autonomy in terms of institutional design or implementation of federal laws. However, this problem is not likely to be important in the countries under consideration, which are rather highly centralized politically. Even in the case of the Russian federation, the strong presidential regime, whose focus has been to recentralize state power (Fish, 2001), helps attenuate the (statistical) problem. In other large countries, such as Poland, Romania, Ukraine, Belarus, and Kazakhstan, power devolution to the regions and the extent of local democracy remain limited.<sup>19</sup>

We thus identify individuals of a given country who live in regions that experience unequal degrees of market development, and we estimate the support for democracy of individual  $i$  living in region  $r$  of country  $j$  on the degree of market development of region  $r$  in country  $j$  ( $\text{Market}_{rj}$ ), controlling for individual sociodemographic characteristics  $X_{ij}$  and for country dummies  $C_j$ :

$$\begin{aligned} \text{SupportDemocracy}_{ij} = & b_0 + b_1 \text{Market}_{rj} \\ & + b_2 X_{ij} + b_3 C_j + u_i. \end{aligned} \quad (2')$$

In order to control for potential omitted cultural factors, we again rely on the long-lasting influence of former empires or federations, and we run equation (2') on the various subsamples of formerly integrated empires and federations and of today's areas of deeper integration.

Admittedly, the identification strategy for the estimation of equation (2') is weaker than that of equation (1''). Indeed, it relies on the assumption that the uneven development of the market across the regions of a country is not due to some local variable that would also affect the political attitudes of the inhabitants. In order to reduce the potential influence of such omitted variables, we control for the type of living area of respondents (metropolitan, urban, or rural areas), or for the geographical distance to the capital city. Of course, this does not entirely solve the problem. For example, one could suspect that the western part of Ukraine is both more prone to democracy and to economic development than the eastern part, because it is influenced by the more modern values and attitudes of its Western neighbors Poland, Slovakia, and Hungary. While we recognize this risk, we assume that the heterogeneity of values and culture inside a given country is smaller than it is on average between different distant countries. The rationale is that there is more cultural distance, on average, between a Lithuanian and a Kazakh than between a Lithuanian and a Pole. This idea is based on the importance of local social interactions in the transmission of cultural traits (Bisin & Verdier, 2001; Glaeser, Sacerdole & Scheinkman 1996). The underlying empirical model is a transposition of gravity models, where cultural proximity is explained instead of commercial intensity. In a companion paper, we estimated a gravity model that allows assessing the role of physical distance and geopolitical factors on cultural attitudes and economic outcomes.<sup>20</sup> Estimating the model on the same LITS data, we found that physical distance indeed increases the dissimilarity in political attitudes and in economic occupation patterns between pairs of locations, while belonging to the same country, contiguous countries, or the same former dynastic empire, decreases it.

#### IV. Results

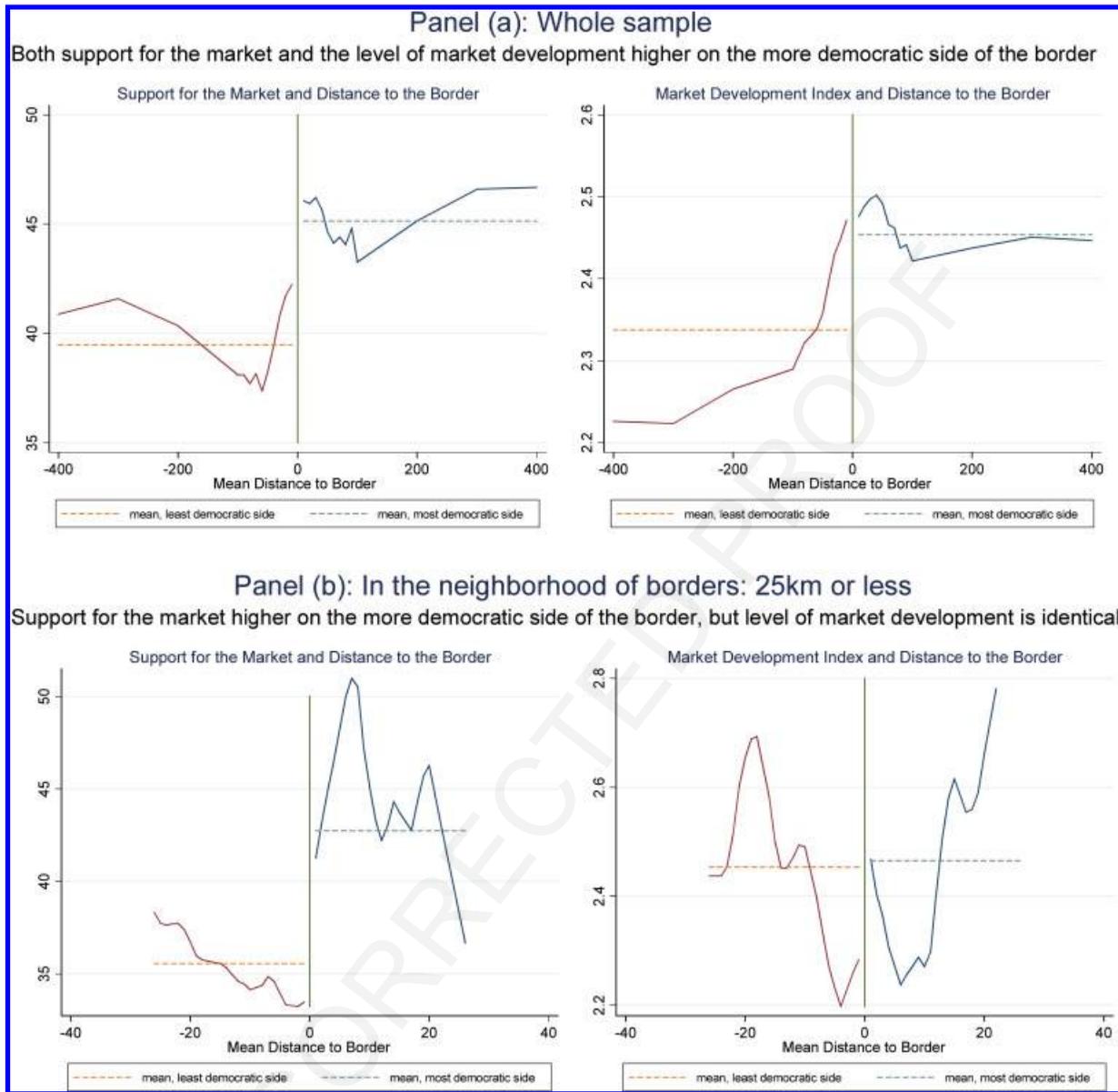
Although simple correlations show that supporting the market and supporting democracy are highly correlated attitudes (the correlation coefficient is 0.45), our identification strategies lead to a more contrasted picture. We find that democracy does increase support for a market economy, whereas the development of the market has no significant effect on the support for democracy.

This pattern is apparent in descriptive statistics. On average, among inhabitants of frontier zones, the percentage of

<sup>19</sup> See, for example, Kovryga and Mooney Nickel (2004) and Riabchuk (2008) for the case of Ukraine, and Regulska (1997) and Yoder (2003) for Poland. For Belarus or other authoritarian regimes in Central Asia, presidential power is an impediment to any form of power devolution.

<sup>20</sup> For a detailed account of the underlying theoretical model and empirical results, see Grosjean (forthcoming).

FIGURE 2.—SUPPORT FOR THE MARKET AND MARKET DEVELOPMENT INDEX AT THE BORDER



(a) Distances are rounded to the nearest 10 for distances up to 100 and to the nearest 50 for distances higher than 100. (b) Distances are rounded to the nearest unit. In both panels, the vertical spike represents the border. Locally weighted scatter plot (lowess) (bandwidth of 0.8) on each side of the border. Horizontal lines are averages on each side of the border.

those who support the market is 7% higher on the more democratic side of the border. By contrast, there is no correlation between regional average support for democracy and the level of market development.

*A. Democracy Increases Support for the Market*

We start with a graphical analysis that illustrates the main assumption of the spatial discontinuity regression. Using longitudinal and latitudinal coordinates of each PSU in the sample, we computed the distance of every PSU in each country to each border of this country.<sup>21</sup> Distances are

<sup>21</sup> We use the great circle formula (Head & Mayer, 2008) in order to compute the physical distance from coordinates.

indexed negatively when one gets closer to the political border on its least democratic side and with a positive sign on its most democratic side (in terms of the Freedom House Nations in Transit ranking). For example, the distance between a PSU in Poland and the Polish-Belarusian border is indexed positively, while the distance between that same PSU and the Polish-Czech border is indexed negatively. We then averaged the measure of support for the market and the index of market development across countries, by ranges of distance to the borders.

Figure 2 presents the percentage of people who support the market and the index of market development as a function of the distance to political borders. Figure 2a includes all PSUs of the sample. Support for the market is clearly

TABLE 1.—DEMOCRACY INCREASES THE SUPPORT FOR MARKET DEVELOPMENT

Dprobit Estimates of Support for the Market	(1)	(2)	(3)	(4)
	Market Preferable	Plan Preferable	Market Preferable	Plan Preferable
Democracy index	0.040*** [0.015]	-0.026 [0.019]		
More Democracy			0.088** [0.034]	-0.067** [0.028]
Adult (35–49)	-0.067*** [0.015]	0.053*** [0.016]	-0.068*** [0.015]	0.054*** [0.016]
Middle age (50–65)	-0.094*** [0.021]	0.055*** [0.021]	-0.095*** [0.021]	0.055*** [0.020]
Old (65 and over)	-0.157*** [0.024]	0.062** [0.031]	-0.162*** [0.025]	0.066** [0.031]
Poor	-0.067*** [0.018]	-0.008 [0.016]	-0.065*** [0.018]	-0.009 [0.015]
Rich	0.016 [0.021]	-0.005 [0.015]	0.011 [0.020]	-0.001 [0.014]
Male	0.051*** [0.010]	-0.001 [0.011]	0.053*** [0.010]	-0.003 [0.011]
Compulsory education	-0.019 [0.032]	0.096*** [0.033]	-0.016 [0.031]	0.100*** [0.031]
Secondary education	0.039 [0.035]	0.131*** [0.034]	0.046 [0.035]	0.130*** [0.032]
Professional education	0.055* [0.032]	0.126*** [0.028]	0.061* [0.032]	0.124*** [0.028]
University education	0.105*** [0.034]	0.193*** [0.034]	0.111*** [0.033]	0.190*** [0.033]
Post graduate education	0.023 [0.076]	0.200*** [0.076]	0.039 [0.078]	0.190** [0.077]
Unemployed	0.000 [0.027]	-0.006 [0.027]	0.006 [0.027]	-0.012 [0.026]
Self-employed	0.096*** [0.027]	-0.088*** [0.018]	0.110*** [0.027]	-0.096*** [0.018]
White-collar worker	0.051 [0.037]	0.006 [0.028]	0.057 [0.036]	-0.004 [0.028]
Blue-collar worker	0.000 [0.033]	0.027 [0.029]	-0.001 [0.032]	0.023 [0.027]
Service worker	0.033 [0.033]	-0.007 [0.025]	0.043 [0.033]	-0.017 [0.024]
Farmer, farmworker	0.031 [0.052]	0.014 [0.033]	0.048 [0.051]	0.001 [0.033]
Pensioner	-0.012 [0.037]	0.004 [0.022]	-0.002 [0.036]	-0.005 [0.023]
Student	0.023 [0.059]	0.011 [0.046]	0.029 [0.057]	0.004 [0.043]
Housewife	0.045 [0.049]	-0.026 [0.028]	0.060 [0.048]	-0.039 [0.028]
Observations	6,690	6,690	6,910	6,910
Pseudo-R <sup>2</sup>	0.0689	0.0331	0.0742	0.0337

Controls: Frontier zone dummies. Omitted categories: young (17–34 years old), middle income group, lowest education, employee, occupation in army. Robust standard errors adjusted for clustering on frontier zones. \*\*\*Significant at 1%, \*\*Significant at 5%, Significant at \*10%. "Democracy index" is built on the basis of Freedom House (2006a). "More democracy" is a dummy variable based on the same index (see section II C).

higher on the "more democratic" side of the borders, but so is the index of market development, so that it is impossible to disentangle the influence of democracy on political preferences from that of market development. Figure 2b retains only PSUs located less than 25 kilometers away from a border. Support for the market is still significantly higher on the "more democratic" side of the border.<sup>22</sup> However, there is no longer any discontinuity at the border in the average level of market development. This indicates that the main

identifying assumption of our spatial regression discontinuity is valid: there is local invariance in the level of market development within 25 kilometers of each border, while there is variation in political institutions and preferences for a market economy.

Figure 2 shows that support for the market is markedly higher on the most democratic side of the border. This effect is confirmed in a regression analysis. Table 1 presents the general regressions of the support for the market on scores of democratization (equation 1'). Columns 1 and 2 present the regression on the Freedom House Nations in Transit (2006a) democracy score, whereas columns 3 and 4 contain the regressions on a dummy variable representing

<sup>22</sup> The  $p$ -value for a negative difference between the means of support for the market between the least and the most democratic side of the border is 0.079.

TABLE 2.—DEMOCRACY AND SUPPORT FOR THE MARKET WITHIN CULTURAL AREAS

	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)
Dprobit Estimates of Support for the Market	Habsburg Empire	Habsburg Empire <sup>a</sup>	Ottoman Empire	Prussia	Polish Lithuanian Commonwealth	USSR	Yugoslavia	CIS	Central Asia
Democracy index	-0.003	0.078***	0.105**	0.172***	0.032**	0.064***	0.085***	0.060*	0.231***
	[0.065]	[0.011]	[0.049]	[0.020]	[0.013]	[0.020]	[0.022]	0.031]	[0.043]
Pseudo-R <sup>2</sup>	0.07	0.08	0.07	0.51	0.11	0.09	0.06	0.10	0.13
Observations	1,976	1,676	1,840	57	1,714	2,914	2,134	2,314	7,40

<sup>a</sup>Without Hungary. Subsamples of frontier zones. Controls: income categories, age categories, gender, occupation categories, self-employed, education categories, and frontier zone dummies. Robust standard errors adjusted for clustering on frontier zones. \*\*\*Significant at 1%, \*\*Significant at 5%, \*Significant at 10%. The democracy index is built on the basis of Freedom House (2006a).

the relative advancement of democracy (see section IIC). All regressions are performed on the subsample of respondents living in frontier zones, less than 25 kilometer away from the border. We control for frontier zone dummies and standard errors are adjusted for clustering on frontier zones. In table 2, the regressions are performed within each zone of deeper cultural integration, as discussed in section III B.

Columns 1 and 3 in table 1 analyze the determinants of the probability of declaring that a market economy is preferable to any other form of economic system. The coefficients on both the democratic index and the “More Democracy” dummy variable are positive and significant. Column 1 displays the marginal effect of a change in the Freedom House democratic score, while the coefficient on the variable “More Democracy” represents the effect of a discrete change of this dummy variable from 0 to 1 (on the probability to support the market). Hence, column 3 in table 1 shows that conditionally on living in a frontier zone, living on the “more democratic” side of the frontier increases the probability of supporting the market by about 8.8 percentage points.

Columns 2 and 4 analyze the determinants of the probability to declare that under some circumstances, a planned economy may be preferable to a market economy. Column 2 shows that an incremental change in the democracy score does not significantly affect the preference for a planned economy. However, the coefficient on the discrete “more democracy” dummy variable is significant at 5%. Column 4 thus indicates that conditionally on living in a frontier zone, experiencing a more democratic regime reduces the probability of favoring a planned economy by roughly 6.7 percentage points.<sup>23</sup>

Other rows of table 1 display the other correlates of attitudes toward the market. We distinguish three income categories (the richest, middle, and poorest quintiles of income inside each country), six educational levels, occupational categories, and employment status (self-employed versus employees). Self-employed workers tend to be more sup-

portive of the market, while elder people and the poorest third of the population are less so.

If one accepts the assumption that people living in a common frontier zone share the same practical experience of the market (and the same background culture), the lesson of table 1 is that living in a country with a higher degree of democracy exerts a positive influence on the support for the market.

As a robustness check, we ran the same regression as in column 3 of table 1 within each frontier zone.<sup>24</sup> The positive effect of democratic institutions on the support for the market was unchanged and proved particularly strong and significant at borders that are well integrated both culturally and economically, for example, the frontier between Moldova and Romania or between Estonia and Latvia. The effect was also particularly strong for the Belarus-Lithuania, Belarus-Poland, and Ukraine-Russia frontier zones, all formerly part of the Polish-Lithuanian Commonwealth and currently highly integrated.

We also estimated equation (1'') within various sub samples of frontier zones belonging to historically integrated regions. Table 2 presents the regressions within cultural zones as defined in section III A. A higher level of democracy exerts a positive and significant influence on the support for the market in the regions of the former Ottoman Empire; Prussia; Polish Lithuanian commonwealth; and countries of the former USSR, the former Yugoslavia, the CIS, and Central Asia. By contrast, it is not significant in the former Habsburg Empire. Nevertheless, when Hungary is excluded from the sample, the effect is positive and significant in the remaining regions of the former Habsburg Empire.<sup>25</sup>

As a robustness check, we also tested different indicators of democracy: the *Freedom in the World* (Freedom House, 2006b), *BTI* (Bertelsmann Stiftung, 2005) and *Polity IV* (CIDCM, 2006) indicators. The result concerning the influence of democracy on the support for the market was preserved.

<sup>23</sup> In alternative specifications not reported in this paper, we verified that our results are robust to alternative specifications, such as OLS or multinomial logit. We also checked that the results are robust to clustering at the country level. Our result holds when people who declare that “it does not matter” are excluded from the sample.

<sup>24</sup> Results are not reported here but are available from the authors on request.

<sup>25</sup> The unexpected results for Hungary may be explained by the difficult situation in the country, which, at the time of the LITS, was going through a political crisis.

We have checked that the relation was robust to the inclusion of other country-level variables, which might have an impact on support for the market and at the same time be correlated with democracy: life expectancy, GDP per capita, the unemployment rate, and the level of government expenditure as a percentage of GDP (a proxy for the importance of the welfare system). These variables were never significant in any of the regressions (in the subsample of border zones).

We also considered the possibility that there might be elements of market liberalization of a country that can be taken advantage of only by citizens of that country. To get around this issue, we controlled for some measures of market liberalization, which could be associated with restrictions to entry and unequal treatment of agents depending on their nationality. We successively included in the regression the EBRD transition indicator of small-scale privatization (which also includes the tradability of land), the EBRD indicator of banking reform, and the EBRD indicator of interest rate liberalization. This did not alter the results. Furthermore, these variables turned out not to be significant, consistent with the idea that there is no variation in the perception of market development across borders in frontier zones.

#### B. Market Liberalization Does Not Raise the Support for Democracy

We now address the symmetric question whether market liberalization increases the support for democracy. As explained in section III, we rely on the evidence that there are wide regional differences within the countries under study.

Table 3 displays various estimates of political attitudes regressed on indices of regional market development using the entire sample—all regions of a given country. Because the impact of market development on political attitudes could be driven by metropolitan regions in which market liberalization is more advanced and where people are likely to have specific political attitudes, we include a control for the type of area (metropolitan, urban, or rural) in all regressions.

Column 1 analyzes the determinants of the probability to declare that democracy is preferable to any other form of political system. The index of market development has no impact on this variable. Identically, the probability of declaring that “under some circumstances, an authoritarian government may be preferable to a democratic one” (column 2) does not depend on the index of market development.<sup>26</sup> Hence, market liberalization does not appear to reinforce democratic values. Other effects indicate that the richer, better-educated, younger, self-employed people, and, surprisingly, farmers and farmworkers are more sup-

<sup>26</sup> In alternative specifications not reported in this paper, we verified that our results are robust to alternative specifications, such as OLS or multinomial logit. The same result holds when people who declare that it “does not matter” are excluded.

TABLE 3.—SUPPORT FOR DEMOCRACY AND REGIONAL INDICES OF MARKET LIBERALIZATION

	(1)	(2)
Dprobit estimates of Support for Democracy/Authoritarianism	Democracy Preferable	Authoritarian Government Preferable
Market development index of region	−0.011 [0.030]	0.002 [0.019]
Adult (35–49)	−0.026** [0.011]	0.018** [0.008]
Middle age (50–65)	−0.032** [0.016]	0.022** [0.010]
Old (65 and over)	−0.061*** [0.023]	0.026* [0.015]
Poor	−0.053*** [0.010]	−0.001 [0.006]
Rich	0.031*** [0.010]	−0.003 [0.005]
Male	0.037*** [0.008]	0.002 [0.005]
Compulsory education	0.044** [0.021]	0.029* [0.017]
Secondary education	0.105*** [0.022]	0.040*** [0.015]
Professional education	0.119*** [0.019]	0.046*** [0.015]
University education	0.188*** [0.020]	0.056*** [0.016]
Postgraduate education	0.254*** [0.026]	0.016 [0.026]
Unemployed	0.023 [0.018]	−0.013 [0.013]
Self-employed	0.032 [0.022]	−0.017 [0.014]
White-collar worker	0.091*** [0.013]	−0.012 [0.011]
Blue-collar worker	0.031* [0.016]	−0.01 [0.012]
Service worker	0.054*** [0.014]	−0.007 [0.011]
Farmer, farmworker	0.066*** [0.019]	−0.001 [0.013]
Pensioner	0.001 [0.019]	−0.005 [0.010]
Student	0.134*** [0.029]	−0.036** [0.014]
Housewife	0.045** [0.019]	−0.037*** [0.014]
Observations	27,920	27,920
Pseudo- $R^2$	0.054	0.027

Controls: country dummies, urban/rural or metropolitan area. Omitted categories: young (17–years old), middle-income group, occupation in army, self-employed, lowest education. Robust standard errors adjusted for clustering at the country level. \*\*\*Significant at 1%, \*\*Significant at 5%, \*Significant at 10%. The market development index is constructed at the regional level.

portive of democracy. By contrast, the poor, those who have not completed compulsory education, and women are less supportive of democracy and more likely to declare that the political system does not matter to them. These results are similar to those of Fidrmuc (2000).

We also ran the same estimation of support for democracy within each country of the survey.<sup>27</sup> Country-wise regressions corroborated the finding that support for democ-

<sup>27</sup> Results are not reported here but are available from the authors upon request.

TABLE 4.—SUPPORT FOR DEMOCRACY AND REGIONAL INDICES OF MARKET LIBERALIZATION IN FORMER EMPIRES AND FEDERATIONS

	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)
Dprobit Estimates of Support for Democracy	Habsburg Empire	Habsburg Empire <sup>a</sup>	Ottoman Empire	Prussia	Polish Lithuanian Commonwealth	USSR	Yugoslavia	CIS	Central Asia
Market development index	0.119*** [0.046]	0.089 [0.092]	0.014 [0.048]	-0.022 [0.128]	-0.004 [0.044]	-0.080** [0.034]	0.045 [0.044]	-0.077 [0.048]	-0.204** [0.096]
Pseudo-R <sup>2</sup>	0.06	0.06	0.06	0.09	0.06	0.05	0.06	0.06	0.05
Observations	4,635	5,635	7,954	239	4,972	13,952	5,971	10,972	4,000

Controls: income categories, age categories, education categories, gender, occupation categories, self-employed, country dummies. Regions (in columns) are defined as in table 2. Robust standard errors adjusted for clustering at country level. \*\*\*Significant at 1%, \*\*Significant at 5%, \*Significant at 10%. <sup>a</sup>Without Hungary. The market development index is constructed at the regional level. Finally, as robustness checks, we use alternative indicators of the explained variable (support for democracy) and of the explanatory variable (market development).

TABLE 5.—SUPPORT FOR DEMOCRACY AND RELATIVE REGIONAL INCOME

	(1)	(2)
Dprobit estimates of Support to Democracy	Democracy Preferable	Authoritarian Government Preferable
Relative regional level of expenditure	0.021 [0.043]	-0.002 [0.031]
Observations	27,960	27,960
Pseudo-R <sup>2</sup>	0.055	0.027

Controls: income categories, age categories, education categories, gender, occupation categories, self-employed, country dummies. Robust standard errors adjusted for clustering at the country level. \*\*\*Significant at 1%, \*\*Significant at 5%, \*Significant at 10%.

racy does not increase with the market development indicator.<sup>28</sup>

In table 4, we estimate the support for democracy within each zone of deeper cultural integration. Essentially regional market development again appears to have no significant impact on the support for democracy. The index of market development is significant only for the regions of the former Habsburg Empire. Again, when Hungary is excluded from the sample, the effect is no longer significant in the remaining regions of the former Habsburg Empire. In countries of the former USSR and of Central Asia, the impact of the market development index is significantly negative.

First, we use the score of “attachment to democracy,” as described in section IIA. We regress this score on the regional market development index, on the whole sample, including the usual controls and clustering at the country level. Again, our regional index of market development does not significantly influence the importance that citizens attach to democratic values. The coefficient estimate is 0.048, and the standard deviation is 0.090.

We then consider the risk that the indices of market development that we use are misconstrued. We use an indicator of relative wealth as an alternative to the market development index. We compute the average aggregate regional income based on the real expenditures declared by

the households of the survey, relative to the national average.<sup>29</sup> This is based on the idea that aggregate income is an outcome of market development. This relative wealth index is thus more specifically an indicator of the successful outcome of market development. Again, as shown by table 5, this indicator does not have any significant impact on the attitudes to democracy or authoritarian regimes. This contrasts with the relationship between average income and democracy that has been observed at the aggregate national level by Barro (1996) or Minier (2001).

Admittedly, the identification strategy is weaker for the estimation of the relationship running from market development to the demand for democracy. This is essentially because of the risk of regional-level omitted variable. In other words, there may be differences in the degree of development of democracy at the regional level, which are correlated with the regional measure of market liberalization. However, these would likely be upward biases, as more democratic countries are also those in which residents are the more supportive of democracy. The fact that we do not observe any statistical impact of the level of market development on the demand for democracy suggests that the bias is actually negligible. Still, in order to address this potential problem, we ran several robustness tests. First, in the main regression (table 3), we controlled for the type of area where respondents live—urban, metropolitan, or rural areas—as different levels of urban concentration are typically correlated with the type of omitted variable that influences the variables of interest (attitudes and advancement in reforms). Alternatively, for the same reason, we controlled for the distance of respondents’ location to the capital city of each country, or dropped metropolitan areas from the sample. This did not alter the results.

In summary, the degree of market development does not seem to exert a sizable impact on the political support for democracy or on the rejection of authoritarian regimes. In contradiction to a frequently held view that democracy will follow market liberalization, developing market institutions is not a guarantee or a sufficient condition of the subsequent

<sup>28</sup> As a robustness test, we ran the same regression as in table 3 on the Moldovan subsample, including the market development index from the Moldovan firm census (see section IID) instead of the LITS-based indicator. The coefficient on the index of market liberalization was never significant in any specification.

<sup>29</sup> Household expenditures were adjusted for household size using the modified OECD equivalence scale.

emergence of demand for democracy, even when the market is successful in raising aggregate income.

#### IV. Conclusion

The main result of this paper is that in transition countries, building democratic institutions has acted as an ingredient in favor of market liberalization, whereas early market development is no guarantee of subsequent popular support for democracy. This contrasts with the widespread view that the demand for democracy needs naturally emerge as a by-product of capitalism. The sequencing literature (Giavazzi & Tabellini, 2005) claims that liberalizing the economy in the first place is best for economic growth. Our results suggest that at a certain point, democratization may become a necessary condition to obtain public support for further liberalization. To be sure, this does not exclude the risk that democracy complicates the task of reformers, with the risk of impeding market liberalization.

A possible explanation for the positive relationship between democracy and support for a market economy is the impact of democracy on the extent of the social cohesion policy implemented by the government. As advocated by Acemoglu and Robinson (2000) and Acemoglu, Bautista et al., (2007), extending democratic rights can play the role of a credible commitment of the government to future income redistribution, thereby contributing to reinforce popular adhesion to the process of economic reform. More generally, citizens of stronger democracies may have more confidence in their ability to obtain a high degree of social insurance in the face of economic shocks.

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## APPENDIX

TABLE A1.—DESCRIPTIVE STATISTICS

Variable		Mean	S.d.
Market preferable	1 if respondent prefers market economy to other form of economic system	0.43	0.50
Planned economy is preferable	1 if respondent prefers planned economy under certain circumstances	0.26	0.44
Democracy preferable	1 if respondent prefers democracy to other form of political system	0.57	0.49
Authoritarian regulation preferable	1 if respondent prefers authoritarian regime under certain circumstances	0.16	0.37
Old	More than 65 years old	0.16	0.37
Middle age	Between 50 and 65 years old	0.24	0.43
Adult	Between 35 and 49 years old	0.31	0.46
Gender	1 if male	0.48	0.50
No education	1 if has not completed compulsory education	0.05	0.21
Compulsory education	1 if compulsory education highest level of education	0.18	0.38
Secondary education	1 if secondary education highest level of education	0.28	0.45
Professional training	1 if professional training highest level of education	0.31	0.46
University	1 if university highest level of education	0.19	0.39
Post graduate	1 if postgraduate studies highest level of education	0.08	0.09
Unemployed	Actively looking for a job. waiting for an answer or find no job available	0.09	0.29
White-collar worker <sup>a</sup>		0.17	0.38
Blue-collar worker <sup>a</sup>		0.18	0.38
Service worker <sup>a</sup>		0.12	0.32
Farmer or farmworker <sup>a</sup>		0.05	0.22
Pensioner		0.21	0.41
Student		0.03	0.16
Housewife		0.06	0.25
Self-Employed	Work as self-employed at their main job (regardless occupation)	0.08	0.28
Democracy index	2006 Freedom House Nations in Transit Democracy Index rescaled such that a higher score reflects a higher level of democracy. <sup>b</sup> Min: 1.18 (Uzbekistan), Max: 6.25 (Slovenia).	4.01	1.61
Market development index	Regional index: sum of shares of active population employed in: SMEs. private. post-1989 created enterprises. Minimum: 1; Maximum: 5.	2.39	1.05
Importance of democracy	Index of importance of above democratic institutions (see section 5). Min: 0; Max: 9.	5.38	3.59

<sup>a</sup>The different categories of employment were constructed from the responses about the respondent's first job, using the ISCO classification. The ISCO categories corresponding to our white-collar category are (1) legislator, senior official, and manager; (2) professionals; and (3) technicians and associated professionals. Our service workers category consists of (4) clerks and (5) service workers and shop and market sales workers. (6) Skilled agricultural and fishery workers, are in our "farmer and farmworker category" together with independent farmers. All remaining ISCO categories (except armed forces, our excluded category) are considered blue-collar workers. <sup>b</sup>The democracy score ratings from Nations in Transit survey by Freedom House are based on a scale of 1 to 7, with 1 representing the highest level of democratic progress and 7 the lowest. It averages ratings for the electoral process, civil society, independent media, independence of the judicial system, and corruption.

TABLE A2.—CORRELATION BETWEEN MARKET DEVELOPMENT INDICES

	Adjacent Frontiers	Adjacent Countries
Whole sample	0.68	0.21
CIS	0.21	-0.08
Central Asia	0.78	0.11
Baltic states	0.90	0.38
European Union	0.73	0.26
Former USSR	0.73	0.27
Former Yugoslavia	0.08	-0.27
Former Ottoman Empire	0.10	-0.18
Former Habsburg Empire	0.10	0.07
Former Polish Lithuanian Commonwealth	0.51	-0.13

Number of observations (frontier zones): whole sample: 65; CIS (Armenia, Azerbaijan, Belarus, Georgia, Kazakhstan, Kyrgyz Republic, Moldova, Russia, Ukraine and Tajikistan): 23; Central Asia (Kazakhstan, Kyrgyz Republic and Tajikistan): 5; Baltic states: 7; European Union: 26; Former USSR (CIS and Baltic States): 31; Former Yugoslavia (Bosnia and Herzegovina, Croatia, FYROM, Montenegro, Serbia and Slovenia): 12; Former Ottoman Empire (see below): 13; Former Habsburg Empire (see below): 15; Former Polish Lithuanian Commonwealth (see Section III A): 20. For each pair of frontier zones between two countries *i* and *j*, the average indices of market liberalization (see section III D.) are calculated at the level of frontier zone *i* and frontier zone *j* and of country *i* and country *j*.

TABLE A3.—RESPONDENTS CHARACTERISTICS ON EACH SIDE OF THE FRONTIER

	On Less Democratic Side <sup>a</sup>	On More Democratic Side <sup>a</sup>	Difference <sup>b</sup>
Age	46.5 [0.293]	48.6 [0.319]	-2.100***
Education	2.46 [0.019]	2.38 [0.021]	0.071*
Proportion of unemployed <sup>c</sup>	0.095 [0.018]	0.077 [0.011]	0.018
Proportion of self-employed <sup>c</sup>	0.057 [0.007]	0.070 [0.018]	-0.013
Proportion of white-collar <sup>c</sup>	0.180 [0.023]	0.137 [0.015]	0.043
Proportion of blue-collar <sup>c</sup>	0.153 [0.012]	0.164 [0.012]	-0.010
Proportion of service workers <sup>c</sup>	0.115 [0.010]	0.111 [0.010]	0.113
Proportion of farmers or farmworkers <sup>c</sup>	0.043 [0.010]	0.055 [0.024]	-0.012
Observations	3,801	3,121	

<sup>a</sup>Mean with standard errors in brackets. <sup>b</sup>Statistical significance of the difference between means across the border: \*\*\*Significant at 1%, \*Significant at 10%. <sup>c</sup>Proportions are computed over the active population. We checked that the proportion of people employed in each category of the two-digit ISIC classification of industries is identical on either side of the border (results not reported).

FIGURE A1.—MAP OF DYNASTIC EMPIRES IN CENTRAL AND EASTERN EUROPE



The figure indicates PSUs that belonged to the Russian, Prussian, Habsburg and Ottoman Empires for more than 200 years. Source: EurAtlas, *Periodical Historical Atlas of Europe* (2008).

UNCORRECTED